

# Human Development Research Paper 2009/44 **Revisiting the Migration-Development Nexus: A Gravity Model Approach**

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# Abstract

This paper presents empirical estimates of a gravity model of bilateral migration that properly accounts for non-linearities and tackles causality issues through an instrumental variables approach. In contrast to the existing literature, which is limited to OECD data, we have estimated our model using a matrix of bilateral migration stocks for 127 countries. We find that the inverted-U relationship between income at origin and migration found by other authors survives the more demanding bilateral specification but does not survive both instrumentation and introduction of controls for the geographical and cultural proximity between country pairs. We also evaluate the effect of migration on origin and destination country income using the geographically determined component of migration as a source of exogenous variation and fail to find a significant effect of migration on origin or destination income.

Key words: Gravity models, international migration, economic growth.

JEL-Codes: F16, F22, O15, O19, O57

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## 1. Introduction

How migration affects—and is affected by—development remains one of the most contentious issues in contemporary policy debates. Advocates of greater openness towards immigration argue that international movements of persons contribute to development at home and destination by eliminating differences in marginal products across regions. Opponents, on the other hand, contend that the labor market effects of migration at destination can adversely affect the bases of social cohesion, while the loss of skilled workers at origin can deteriorate the chances of poor countries of sustaining the provision of goods that are basic for development.<sup>1</sup> At the same time, any analysis of the migrationdevelopment nexus is made particularly complex by the fact that we expect migration to be affected by the home and destination country's economic prosperity.

In this paper, we try to evaluate the magnitude and form of these links through an empirical analysis of bilateral migration stocks. Seeking to answer whether and how income and emigration influence one another, we quantitatively analyze the drivers and effects of migration with a particular focus on bidirectional chains of causality using an instrumental variables approach. To study whether migration causes income we use the geographically-induced variations in migration across countries as our source of exogenous variations. Conversely, to study whether income causes migration we use the variation in European settler mortality rates as a source of exogenous variation in levels of development.

A careful identification of the multiple potential links of causation between development and income is relevant not only to academic debate, but also to current policy discussions in both developing and developed countries. Many policymakers in developed countries argue that one—if not the only—way of curbing immigration pressure 'here' is to foster economic development 'there'. This approach depicts immigration as a problem and

<sup>&</sup>lt;sup>1</sup> For the argument in favor of migration, see UNDP (2009), World Bank (2006), and World Bank (2009). For critiques, see Borjas (1999), Chen and Boufford (2005), Mills, et al (2008), and UNCTAD (2007).

lack of development as its root cause.<sup>2</sup> Yet, historical and ongoing international migration patterns point to a more complex picture. For most of the countries that are experiencing a rise in living standards, economic development may well be associated with *higher* emigration rates, at least until a certain development level is achieved.<sup>3</sup> For example, the median emigration rate in countries with low levels of human development is 3.66, in contrast to that in countries with high levels of human development—7.18.<sup>4</sup> Because of these patterns, some historians and migration scholars have hypothesized the existence of an inverted-U relationship between development and emigration based on either cross-sectional or time-series comparisons.<sup>5</sup>

However, simply observing that development and emigration appear to display such a non linear-relationship leaves important questions unanswered. First of all, it is not hard to come up with a story in which such a relationship is generated by reverse causation. Further, even if the relationship reflects causation from development to migration, one must wonder what it is about the process of development that impacts individuals' propensity to leave or stay. It may be rising income, but it may also be other socioeconomic and demographic changes brought about by, contributing to, or simply associated with rising income. Ultimately, people don't just take a decision to migrate. They decide to migrate *to another* country, which is why one should try to take into account the characteristics of receiving and sending countries, as well as of the idiosyncratic forces affecting pairs of countries in any

<sup>&</sup>lt;sup>2</sup> For instance, in 2005, European Commission's President José Manuel Barroso stated, on the occasion of the EC approval of the European Union Strategy for Africa that "the problem of immigration, the dramatic consequences of which we are witnessing, can only be addressed effectively in the long term through an ambitious and coordinated Development cooperation to fight its root causes", and the following year Nicolas Sarkozy, then France's Interior Minister, stated in Rabat, Morocco, that "the development of Africa is the only solution, the only answer to the challenge of immigration".

<sup>&</sup>lt;sup>3</sup>A country's emigration rate in a given year is defined as its emigrant population as a share of its total native born population (at home and abroad) in that particular year.

<sup>&</sup>lt;sup>4</sup>We use the Human Development Report's country groups by levels of the home development, which is indexed by the HDI as follows: low, 0.000-0.499; medium, 0.500-0.799; high, 0.800-0.899; and very high, 0.900-1.000. See UNDP (2009), which also refers to the last group as "developed" countries. Note that high human development countries are *not considered developed countries but are rather the highest HDI group among developing countries*.

<sup>&</sup>lt;sup>5</sup> See, for example, Hatton and Williamson (1994) and de Haas (2007).

complete migration model (e.g., the distance between them, the existence or absence of a common border or language, the nature and strength of their historical ties). To take a simple example, over 11% of Mexicans are living abroad, compared to 2% of Argentines, despite the similar levels of development shared in both countries.<sup>6</sup> The fact that Mexico shares a common border with the United States, while Argentina does not share a common border with any developed country is surely part of the explanation. The impact of such effects will not be captured in any cross-national empirical exercise that lumps together migrants that are going to very different places. Disentangling and estimating the effects of these various factors in driving and shaping emigration at the global level constitutes the first objective of this paper.

Before exposing our strategy and presenting detailed results, it is important to discuss how we can think and what we know about both chains of causation. Let us first turn to the effect of development on the incentive and propensity to out-migrate, focusing on the potential effect of income. Take the example of a young Turkish woman who considers migrating to Germany but ends up accepting a job in a newly-established joint-venture, and consider alternatively the case of a young man in Bamako who after operating a cyber-café for a few years uses his savings to join his relatives in France. The former example is consistent with the idea that economic development reduces the incentive for emigration by expanding opportunities at home. The latter case illustrates the opposite view: economic development increases individuals' financial propensity to emigrate by alleviating liquidity constraints on movement.

These ideas can be represented through a very simple model illustrating the mechanisms that may generate a non-linear relationship between income and development. An individual decides to migrate by comparing utility of income, *y*, at home,  $U(y_o)$ , and at destination,  $U(y_d)$ , with the cost of migrating, *c*. For simplicity, we use a Cobb-Douglas utility function of the form  $U(y) = y^{\alpha}$ ,  $0 < \alpha < 1$ . The migrant is prevented from moving

<sup>&</sup>lt;sup>6</sup> Mexico has an HDI of 0.854 and Argentina has an HDI of 0.866.

if the cost of migration is greater than income at origin, so we must have  $c \leq y_o$ . Suppose also that there is a randomly distributed term  $e \sim U(0,b)$ , which is the net gain in income at destination and will vary across individuals. Let b < c. We also assume for simplicity that all individuals have the same income  $y_o$  at origin. It follows that an individual migrates if  $U(y_d + e) > U(y_o + c)$ , subject to  $c < y_o$ . This result yields two possible outcomes:

1) If  $y_o$  is sufficiently low, then the cost of moving is too great and emigration is zero.

2) If  $y_o \ge c$ , then the fraction of people migrating will be given by the integral

$$\int_{y_{o}}^{y_{d}+e} U(y)dy = \int_{y_{o}}^{y_{d}+e} y^{\alpha}dy$$
  
=  $\frac{1}{1+\alpha}y^{1+\alpha}|_{y_{o}}^{y_{d}+e}$   
=  $\frac{1}{1+\alpha}[(y_{d}+e)^{1+\alpha} - (y_{o})^{1+\alpha}]$  (1)

As  $y_o$  increases, the fraction of people who migrate decreases. For higher income at origin, there must be a greater net gain of migration to incite movement. The migration hump is represented by Eq. (1) and illustrated in Figure 1.

$$f(y_o) = \begin{cases} 0 & \text{if } y_o < c \\ \frac{1}{1+\alpha} [(y_d + e)^{1+\alpha} - (y_o)^{1+\alpha}] & \text{if } y_o \ge c \end{cases}$$
(2)

The preceding discussion illustrates that it is very important to consider nonlinearities in the study of the effect of income on migration. The idea that over some range of the development process development and emigration may go hand-in-hand has been presented by various authors in the past. Hatton and Williamson (1998, 2003) show that many countries that are highly developed<sup>7</sup> today (e.g., Ireland, England, Italy, Spain, Portugal, and the Republic of Korea) experienced rising emigration rates in the past. As these countries grew wealthier, they also became more attractive to migrants from less developed

<sup>&</sup>lt;sup>7</sup> See UNDP (2009).

countries and were ultimately transformed from net emigration to net immigration countries over a few decades.

Other authors have pointed out that this non-linear relationship also characterizes the cross-sectional data.<sup>8</sup> In fact, around 2000, the emigration rate of Morocco, a middle income country with per capita GDP (PPP) of roughly \$4,000, was twice as high (8.1%) as Niger's 4.0% (with a per capita GDP of less than \$2,000) and Norway's (3.9%), which has one of the world's highest GDP per capita (close to \$50,000 in 2007). Using Human Development Index (HDI) categories yields a similar picture (Figure 2).

Anticipating our strategy and results, we can also point to the fact that in our OLS regressions, the coefficients on origin income and origin income squared retain their significance even when adding other variables, and that their respective signs (positive for the former, negative for the latter) do confirm the existence of an inverted-U relationship between income at origin and emigration. Overall, these facts strongly suggest that economic development and (e)migration do tend to go hand-in-hand—i.e. are positively correlated—at least up to a certain level of development, rather than what is typically posited in policy circles. The results of our OLS regression (1) indicate that 75% of countries in our sample are located on the upward sloping portion of the hump.

But the model that we sketched above suggests that comparisons that are only centered on origin countries may be missing a fundamental part of the picture. In fact, equation (2) predicts that both income at origin and destination countries matter and that they matter nonlinearly. This suggests that models of migration that use either (i) a crosssectional analysis with countries as the unit of observation and thus do not distinguish between migration among different country-pairs, or (ii) a simple linear relationship where migration either increases or decreases with income levels or differentials, may be severely mis-specified. To the best of our knowledge, our paper is the first attempt to address these two specification issues simultaneously.

<sup>&</sup>lt;sup>8</sup> See, in particular, de Haas (2007).

A third issue that arises in the analysis of how migration evolves with development is that we would also expect migration to have an effect on economic efficiency and income. These effects may differ by level of development. For example, emigration may foster economic growth among poorer countries where it can alleviate labor market pressures while providing much-needed revenue and foreign-exchange via remittances. However, in countries with higher levels of income and overall development (including better education systems), high rates of emigration could result in the loss of human capital that is necessary for the development of high-technology industries and overall further development progress. Such a pattern would generate an inverted-U relationship between migration and income that would nevertheless not reflect causation from income to migration.

This paper attempts to deal with the need to understand the migration-development nexus through an analysis of bilateral migration stocks that takes nonlinearities in origin and destination income seriously while simultaneously tackling issues of causality through the use of convincing instruments. This is achieved by estimating a model of bilateral migration stocks using the World Bank/Sussex Database of Bilateral Migrant Stocks<sup>9</sup> (hereafter the Sussex matrix) from which we extract information on bilateral migrant stocks for 127 countries. We deal with the endogeneity of income to emigration using European Settler Mortality (ESM) rates as instruments for income to isolate its exogenous component and potential causal effect on emigration. Our research builds on the work of Acemoglu, Johnson, and Robinson (2001) (hereafter, AJR), who have shown that ESM rates explain contemporary differences in development through their effect on past and present-day institutions. Subsequently, we deal with the endogeneity of emigration by constructing the geographically-determined fraction of bilateral migrant stocks from our gravity model of bilateral stocks and using them as instruments to estimate the effect of migration on income both at origin and at destination for the 127 countries in our sample.

<sup>&</sup>lt;sup>9</sup> Migration DRC (2007).

Our key findings can be summarized as follows. First, we confirm in the model of bilateral flows the existence of a robust non-linear inverted-U relationship between income at origin and emigration. Thus the migration 'hump' pattern exists in the more demanding empirical specification of bilateral migration stocks. Second, we show that the relationship survives in the simplest of our instrumental variables specifications. The relation appears to be particularly fragile to controlling for destination country income and variables capturing the economic, historical and cultural proximity between source and destination countries. Third, when we study the effect of migration on income we do not find a robust effect on either origin or destination income. This supports the idea that migration is not an important contributor nor hindrance to development and that may be best seen in terms of the expanded opportunities that it offers *individuals* to carry out their life plans.<sup>10</sup>

The rest of this paper is structured as follows. Section two reviews the current state of knowledge on international migration patterns, especially those works that have focused on the migration-development nexus and/or used gravity models of migration. In section three, we present our strategy and data. Our results are exposed and discussed in section four. Section five concludes.

## 2. Theoretical and Literature Review

The migration-development nexus has received considerable attention in the literature. This section focuses on two sets of studies in the field: those that have analyzed the empirical relationship between income and development using national-level data, and those that have used gravity models to analyze the determinants and/or effects of migration—often understood as *im*migration. Our paper is at their intersection of these two literatures.

A considerable body of work has dealt with the non-linear relationship that characterizes the effect of development on migration. The term 'migration hump' seems to have been coined by Martin (1993) when discussing the likely effects of NAFTA (North

<sup>&</sup>lt;sup>10</sup> See UNDP (2009), Chapter 1.

American Free Trade Agreement) on irregular migration from Mexico to the United States. Martin argued that NAFTA would stimulate migration in the short to medium run by fostering labor supply and mobility—especially from rural areas—before eventually reducing the incentives to out-migrate as the income gap narrowed with the United States. Martin and Taylor (1996) further argued that the process of social and economic development in its broadest sense tended to be associated with generally higher levels of mobility by helping would-be migrants pay for the fixed costs associated with migration. Only after a longer period of sustained growth and decreasing wage gaps between origin and destination countries would labor migration tend to decrease. According to them, emigration would tend to decrease steeply when the income ratio between receiving and sending countries declines from 4 or 5 to 1. Increasing income inequality would also increase people's incentives to migrate abroad even if average income increases.

These authors also underlined the fact that the downward-sloping portion of the migration hump was by no means inevitable if economic growth did not result in significant employment opportunities, in which case it could result in a semi-permanent 'migration plateau' of sustained out-migration that could last for an undetermined period of time. Olesen (2002) introduced the notion of a migration 'band' to refer to the income range associated with the highest emigration rates between two countries. Olesen (2002) posited that bilateral migration should peak and then decrease when the income differential between the sending and the receiving countries reaches and subsequently falls below a ratio of between 3 and 4.5 to 1, which he terms the 'migration turning point'. More recently, de Haas (2007) discusses how policies of rich countries that seek to stem migration by helping foster development are ill-founded. As the poorest are empowered through economic and human development, they will tend to move to more developed countries to realize even greater gains, thus increasing migration rates over the short and medium terms. Hatton and Williamson (1994, 2002, 2003, 2004, 2005) and Massey (1988, 1990, 2003) have also pointed to historical evidence illustrating the same key fact: at low levels of income, development seems to foster emigration.

Frequently applied in the study of international trade, gravity models have also been applied to analyze the drivers of migration and how, in turn, migration affects income. As the name suggests, gravity models are loosely derived from Newton's law of gravitational force and posit that the interaction between two geographic entities, through trade or migration, are subject to forces that are inversely proportional to the distance (or income differential) between them and on some relevant measure of their 'masses', including population, area and/or income. Gravity models also typically include other geographical controls, such as whether the country is landlocked and distance to the equator, as well as bilateral controls that capture 'pair-specific' characteristics (e.g., whether two countries share a common border, a common language, and a colonial past). The central premise behind these models is that these structural features are likely to determine a country's international trade and migration patterns.

The application of gravity models to the analysis of trade goes back to Tinbergen (1962) and Linneman (1966). Since then, gravity models have been widely used by researchers, including Anderson (1979), Bergstrand (1985, 1989), Deardorff (1995), Frankel and Romer (1999), Egger (2000), and Carrillo and Li (2002). Under quite similar assumptions on the effects of geographical and other structural factors on population movement, gravity models have also been used to study international migration.<sup>11</sup> The overwhelming majority of these studies have concentrated on OECD countries and used data on flows. Lewer and Van den Berg (2008), for instance, estimated a series of gravity equations using panel data for 16 OECD countries for the period 1991–2000. Their regressions estimated bilateral flows between countries, controlling for their populations, the distance between them, the ratio of their per capita incomes, the pre-existing stock of foreign-born migrants, common language, geographical contiguity (common border) and colonial ties. Other controls were also added

<sup>&</sup>lt;sup>11</sup> According to Lewer and Van den Berg (2008) for instance, "Immigration, like international trade, is driven by the attractive force between immigrant source and destination countries and impeded by the costs of moving from one country to another", which "(...) are likely to be correlated with the physical distance between countries", while "(...) ceteris paribus, the more people there are in a source country, the more people are likely to migrate, and the larger the population in the destination country, the larger is the labor market for immigrants."

in different specifications, such as variables for human capital and the rule of law. Their results confirmed that international migration is indeed subject to and driven by 'gravitational-like' forces.

More recently, Mayda (2008) and Ortega and Peri (2009) have estimated gravity models of international migration to developed countries. While their objectives and specifications differ, both papers focus on international migration flows to a subset of OECD countries for the period 1980-1995 (Mayda, 2008) and 1980-2005 (Ortega and Peri, 2009). The authors also use panel data drawn from the OECD's International Migration Statistics based on OECD's Continuous Reporting System on Migration (SOPEMI).

Mayda's results are broadly consistent with the main theoretical predictions of international migration theories. According to her findings, immigration flows are positively correlated to the destination countries' GDP per worker. However, the effect of GDP per worker is not found to be statistically significant. These contrasting results emphasize the importance of so-called 'pull factors' in driving international migration. They also underline the complex nature of the relationship between origin country GDP and emigration, as low income constitutes both an incentive and an impediment to movement.

Geographic and demographic factors also appear to play a major role. Hatton and Williamson (1998, 2003) and Mayda's (2008) results suggest that the share of the origin country's population aged 15-29 has a significant positive impact on outmigration (as a share of the population). Restrictive immigration policies are found to partly offset the effect of other push and pull factors, as, for instance, the impact of distance is greater when policies are relaxed. These findings suggest that 'migration quotas matter'<sup>12</sup> and that the 'asymmetric effect' between destination and home countries GDP is explained by the positive impact of economic growth at destination on policy stance towards immigration.

A concern in most empirical studies of income is endogeneity. Mayda (2008) deals with the endogeneity of income by relating current emigration rates to lagged values of income and then controlling for the endogeneity of income levels with terms of trade as an instrument. This approach would be valid only if were assumed that terms of trade have no direct effect on migration aside from its effect through lagged income, which would not be the case if migrants go to work in significant numbers to tradable-producing industries which are made more profitable by a terms of trade shock.

All of these papers concentrate on the migration-development nexus in OECD countries. To the best of our knowledge, there are no studies analyzing the effect of income on migration using bilateral migration flows which include migration to developing countries, and this paper is the first attempt to study the effect of development on migration using data on migrants from and to both developing and developed countries. This is striking because migration to non-OECD countries accounts for 51 percent of international migration and for 65 percent of all international migration originating in non-OECD countries.<sup>13</sup>

To the best of our knowledge, only two papers have previously used bilateral migration models to study the effect of migration on income. Both use geographical factors as instruments, as we do. Ortega and Peri (2009) use the share of migration explained by geography and demography as an instrument for total migration. However, their study is limited to OECD migration data and only examines the effect of migration on destination income—it is thus silent on the effect of migration on development as such. Felbermayr et al. (2008) use the Sussex matrix of bilateral migration stocks (which we also use) to study the effect of migration on income in both developed and developing countries. Similarly to us, they use an IV technique inspired by Frankel and Romer (1999) to test the effect of immigration on per capita income. They find that a positive and statistically significant effect whereby their preferred specification suggests that a 10% increase in the number of migratis leads to a 2.2% gain in per capita income. Our paper shows, however, that this effect is not robust to adequately controlling for the endogeneity of institutions (Rodrik, et al, 2004).

<sup>&</sup>lt;sup>13</sup> Migration DRC (2007).

## 3. Empirical Strategy

## 3.1. Empirical Model

#### 3.1.1. OLS

In order to test for the migration 'Hump', we use the following empirical specification:

$$ln\Gamma_{ij} = \alpha_0 + \alpha_1 \mathbf{S}_i + \alpha_2 \mathbf{R}_j + \alpha_3 \mathbf{P}_{ij} + \epsilon_{ij}$$
(3)

where  $\Gamma_{ij}$  is the estimated *stock* of individuals born in country *i* living in country *j*; **S** *i* is a vector of country-specific variables of the sending country *i*; **R** *j* is a vector of country-specific variables of the receiving country *j*; **P** *ij* is a vector of pairwise variables between the sending country *i* and receiving country *j*; and  $\epsilon_{ij}$  is an *iid* error term for sending-receiving country pair *ij*. The variables that constitute S, R, and P are discussed below with a full list and sources in Table A1.

#### 3.1.2. Two-stage least squares

To control for the endogeneity of income we use ESM rates and its square. As AJR (2001) demonstrates, ESM rates are correlated with past and present day institutions and thus with present-day development.<sup>14</sup> As a result, they are exogenous to present development and serve as an excellent candidate for an instrument. In order to use ESM as a valid instrument for income, the one condition that we would have to accept is that they do not have an independent effect on migration over and above that which they have on income. This assumption may be questionable in the simplest specifications in which we do not control for current institutions, as in such a case ESM may be correlated with the disturbance term. However, once we control for institutions as well as a host of other country-specific variables—and verify that the instrument maintains explanatory power in the first stage

<sup>&</sup>lt;sup>14</sup> In contrast to AJR (2001), in order to maintain both developed and developing countries in our sample, we set ESM to zero for countries that have never been colonies.

regressions (see Table 2)—the hypothesis of a direct effect of ESM on current migration appears much less tenable. In an alternative specification, we also include former colonial status in the list of instruments. Again, excludability of this instrument may be questioned but is less likely a problem in our most demanding specifications which also include colonycolonizer pair dummies in the explanatory variables of the second stage regression.

## 3.2. The 'Horserace': income and migration

Following a framework similar to that established by Rodrik, Subramanian, and Trebbi (2004), we explore whether or not migration plays as prominent a role in a crosssectional analysis of income as other factors, namely institutions, trade, and geography.

#### 3.2.1. OLS

Using OLS, we estimate the following empirical specification:

$$lny_i = \delta_0 + \delta_1 Inst_i + \delta_2 Trade_i + \delta_3 Geo_i + \delta_4 Mig_i + \epsilon_i$$
(4)

where *i* is the country index; *y* measures income; *Inst*, *Trade*, and *Geo* measure the quality of institutions, the prevalence of trade, and geographical characteristics, respectively. *Mig* represents alternatively immigration or emigration ratios (that is, the ratio of migrants and destination or origin population), which we estimate separately. As above,  $\epsilon_i$  is an *iid* error term.

#### 3.2.2. Two-stage least squares

In Eq. (4), we deal with the likely endogeneity of institutions, trade, and migration using instrumental variables. In keeping with the Rodrik, et al (2004) approach, we use ESM rates as an instrument for institutions and the Frankel-Romer measure of openness for trade. For both immigration and emigration, we use the migration ratios explained by geographical variables, which we know to be exogenous. More precisely, we use the predicted values of regressing the stock of migrants only on the geographical variables of the model:

$$ln\Gamma_{ij} = \eta_0 + \eta_1 Dist_{ij} + \eta_2 Border_{ij} + \eta_3 AbsLat_i + \eta_4 Land_i + \eta_5 Land_j + \epsilon_{ij}$$
(5)

where *Dist* is the distance between origin country *i* and destination country *j*; *Border* is a dummy variable that equals 1 if *i* and *j* share a common border and 0 otherwise; *AbsLat* is the absolute latitude of origin country *i*; and *Land* is a dummy equal to 1 if the respective country is landlocked and 0 otherwise. The predicted values,  $\hat{\Gamma}_{ij}$ , are then summed separately for each origin and destination country,

$$\sum_{j=1}^{k-1} \hat{\Gamma}_{ij} = \hat{\Gamma}_i \text{ and } \sum_{i=1}^{k-1} \hat{\Gamma}_{ij} = \hat{\Gamma}_j,$$
(6)

where k = 181, the total number of countries in our sample.

Thus,  $\hat{\Gamma}_i$  and  $\hat{\Gamma}_j$  represent the geography-predicted components of immigrant and emigrant stocks, respectively. For a single country, i, which has both an immigrant and emigrant stock, we write  $\hat{\Gamma}_i^I$  and  $\hat{\Gamma}_i^E$ , using obvious notation. Dividing those by population, we derive exogenous instruments for immigration and emigration ratios.

Estimating the endogenous variables on all of the exogenous instruments, we have the following first-stage regressions:

$$Inst_{i} = \gamma_{o} + \gamma_{1}ESM_{i} + \gamma_{2}ImG_{i} + \gamma_{3}EmG_{i} + \gamma_{4}FR_{i} + \epsilon_{Inst,i}$$
  

$$Trade_{i} = \theta_{o} + \theta_{1}FR_{i} + \theta_{2}ESM_{i} + \theta_{3}ImGeo_{i} + \theta_{4}EmGeo_{i} + \epsilon_{Trade,i}$$
  

$$Im_{i} = \lambda_{o} + \lambda_{1}ImGeo_{i} + \lambda_{2}EmGeo_{i} + \lambda_{3}ESM_{i} + \lambda_{4}FR_{i} + \epsilon_{Im,i}$$
  

$$Em_{i} = \zeta_{o} + \zeta_{1}EmGeo_{i} + \zeta_{2}ImGeo_{i} + \zeta_{3}ESM_{i} + \zeta_{4}FR_{i} + \epsilon_{Em,i}$$
(7)

where *ESM* are European settler mortality rates; *FR* is the Frankel-Romer index;  $ImGeo = \hat{\Gamma}^{I}/Pop$  and  $EmGeo = \hat{\Gamma}^{E}/Pop$  the exogenous migration rates calculated above; and the  $\epsilon$  's are the respective *iid* error terms.

## 3.3. Testing for the migration 'Hump'

A key feature of our main model is that we estimate stocks—rather than flows—of foreign-born migrants. This characteristic is a direct consequence of the paper's goal of extending the analysis to the whole world rather than restricting it to a small subset of rich countries. As was previously underlined, the majority of recent papers on the subject use the only existing reliable data on flows of international migrants to OECD countries. In contrast, we use the only comprehensive global database-the World Bank/Sussex Bilateral Matrixwhich provides estimates for stocks in or around the year 2000. The inclusion of controls, such as GDP per capita, life expectancy and educational attainment, reflect the widelyviewed notion that absolute and relative levels of income and human development affect migration patterns. Their effects, however, are complex. As discussed earlier, the propensity to emigrate from a low-income country could increase with its level of income and/or development, at least up to a certain point. However, it is also possible that the relationship between the level of development and the propensity to *attract* migrants may be non-linear. When the poorest are able to migrate, they may only be able to move to other very poor countries since the barriers of moderately poor and rich countries are higher. However, once they have sufficiently high income to overcome these barriers, they may move to the wealthiest countries, essentially bypassing the middle income countries. As Klugman and Pereira (2009) show, many developing countries also have high barriers to migration. Once migrants have sufficient income to gain entrance to middle income countries, they may have the means to gain entrance to the wealthiest countries as well. All else equal, migrants would then tend to move to where income gains are greatest.

### 3.4. Data and controls

We investigate how geographic, demographic, cultural, political and economic factors affect the level and composition of the stock of international migrants living in a specific country and groups of countries. The dependent variable is the log of the bilateral stock or flow of migrants between 127 selected countries. The choice of controls rests on an analysis of the current knowledge and literature on potential drivers of international migration (Yang, 2008; Massey, 1990; and Hatton and Williamson, 2002), as well as on the underlying assumptions of gravity models.

#### 3.4.1. Controls

#### Geography

Greater distance between two countries is expected to decrease the propensity of people to move between them, as costs to move tend to increase with distance. We also include dummy variables for countries that share a border and for landlocked countries. Bordering countries often share a unique relationship and tend to be closely linked politically, economically, and culturally. Migrants to or from landlocked countries also face unique constraints since they are unable to travel by water, a common way for people to travel cheaply over long distances.

#### **Demography, Culture and Living Standards**

Demographic factors can also influence migration patterns (Hatton and Williamson, 1998, 2003). Differences in population density and age structure between countries are likely to have a significant impact on bilateral flows. For instance, countries with high fractions of non-working age populations may require more workers from abroad to counteract increasing dependency ratios, and pressure to out-migrate may be greater in countries with a high proportion of young adults.

Countries that share a common language also tend to share other historical and cultural ties. These commonalities can help facilitate a migrant's integration into the host society and labor market. Thus, countries where a potential migrant can already speak the local language become much more attractive destinations. After controlling for income, one could expect health and education to act as 'pull' factors. On the other hand, low health and education levels imply low levels of human capital among natives at destination, which may make migrants more competitive vis-à-vis natives in the labor market.

#### Political stability and governance

Environments characterized by poor governance or political instability may induce people to move to countries where conditions are better. Politically driven migration scenarios may vary from families that flee violent conflict as refugees to those who seek countries where property rights are more secure or basic political freedoms are guaranteed. On the contrary, there are reasons we may also find that higher democracy leads to lower immigration as voters may be more likely than elites to resist immigration, particularly in labor-scarce countries. Moreover, barriers to immigration are more likely to be effective if the country has better institutions to enforce them.<sup>15</sup>

#### **Economy and Trade**

Particular characteristics of a nation's economy may make it a more or less likely to be either a net sending or receiving country. Export controls attempt to estimate whether international migration is a substitute or a complement to trade. The contribution of the service sector to the national economy could be a pull factor, as countries with large service sectors may demand foreign labor. On the other hand, sectors such as services that do not compete internationally may have more organized labor movements and be better able to resist increased immigration.

#### **Measurement Issues**

Data on migration flows is also susceptible to more problems of measurement than stock data. While any cross-country comparison of migration data suffers from differing definitions of who is a migrant (some countries use place of birth while others use nationality), flow data also tends to omit both emigrants and irregular migrants. Stock data typically includes all migrants irrespective of their legal status, while OECD flows data only accounts for legal immigrants. As Ortega and Peri (2009) describe, flow data often poorly

<sup>&</sup>lt;sup>15</sup> In principle, this could be accounted for by controlling for immigration policies, as we do. However, our measure of policies is quite crude. We use an immigration policy index from the UNDESA Population Division's *World Population Policies 2007* (UN, 2008). The index takes only three values: -1 if policymakers support a reduction in immigration levels, 0 to maintain, and 1 to increase.

measures outflows since people who are leaving a country are less likely (or less likely to be required) to report exiting a country.

## 4. Results

## 4.1. The effect of development on migration

Let us first discuss how development affects emigration. Recall that historical and contemporary data tend to show quite consistently that there exists a non-linear, inverted-U shaped relationship between development at origin and emigration.

#### 4.1.1. OLS

Our OLS results confirm the existence of this relationship when origin income is allowed to pick up such non-linearity. The relationship appears to be robust as it continues to hold even after controlling for the effect of various geographic, demographic, historical, and policy variables. As shown in Table 1, the coefficients on origin income and origin income squared remain highly significant throughout regressions (1) to (5) as we progressively control for additional explanatory variables, and their respective signs confirm the inverted-U shape. Furthermore, their magnitude does not appear to be greatly affected by the inclusion of these other controls.

Figure 3 plots the emigration levels for varying income levels as predicted by the coefficients yielded by regressions (1)-(5). Countries with lower income fall on the upward sloping portion of the curve, while those with higher income fall on its downward sloping portion. We estimate that for roughly three-quarters of countries in our sample, higher income is associated with higher emigration. We also estimate that emigration reaches a maximum at an income level of around \$14,500, which is higher than the GDP per capita of countries like Libya and Malaysia in 2007.

To give a sense of the intensity of the relationship, in Table 3 we calculate the standardized coefficients corresponding to Table 1. The standardized coefficients provide an estimation of the magnitude of the change in emigration associated with a one standard

deviation change that a particular independent variable has on migrant stocks.<sup>16</sup> Standardized coefficients were calculated for income separately for countries located on either portion of the curve. The standardized coefficients for the OLS model suggest that among the set of relatively less developed countries (i.e., those located on the upward sloping portion of the hump) a rise in income by one standard deviation in an average sending country is associated with a 27% increase in the number of migrants located in an average destination country. For countries located on the downward portion of the hump, an increase in income is associated with a 6% decrease in emigration.

This finding is interesting in itself. It suggests that if cross-sectional patterns are any indication of future trends, then we should expect that rising income in the vast majority of countries in the world will be associated with higher emigration in the foreseeable future, not the opposite. At this point, however, because of potential endogeneity, we are unable to tell whether this robust relationship between income at origin and emigration reflects a direct causal effect of the former on the latter. As described in Sections 1 and 3, we subsequently used ESM rates as instruments for income (and their squared terms for the squared income terms) in our regressions in order to address endogeneity concerns.

### 4.1.2. Two-stage least squares

#### **First stage**

We first need to confirm that our instruments are valid. Following AJR (2001), we are confident that ESM rates are largely exogenous—in any event today's levels of development and income cannot influence them. Subsequently, we need to check that they are correlated with income today. Table 2 reports the results of various variations of our first stage regressions. These results confirm that each of our four instrumental variables is strongly

 $<sup>^{16}</sup>$  We calculate the standardized coefficients of income for countries above and below each threshold as follows: sdi \* [b(inci) + [ 2\*b(inc<sup>2</sup>i) \* meani]]

where *sd* is the standard deviation of income (measured as the natural log of GDP per capita) below/above the threshold, b(inc) and  $b(inc^2)$  are the OLS coefficients of income and income-squared, respectively, *mean* is the average of income below/above the threshold, and *i* denotes either origin or destination. Thus we have four standardized coefficients for income: origin-above, origin-below, destination-above, and destination-below.

correlated with the potentially endogenous income variable that it is used as an instrument for. Importantly, the statistical significance of the correlation is not subject to the inclusion or exclusion of institutional controls. In other words, our instruments affect income through other channels and mechanisms rather than present institutions. This is critical because, as we will see, the origin and destination 'Rule of law' variables drop out of our preferred IV specification, where we have retained only significant explanatory variables not present in the baseline specification. The data therefore tell us that present institutions do not have a separate effect on migration over and above that of current (instrumented) income. Finally, we are confident that we control for a sufficiently large set of variables that may impact emigration, so that the condition that our instruments be uncorrelated with the disturbance term is also likely met.

#### Second stage regressions

The results of our 2SLS regressions are reported in panel 2 of Table 1.<sup>17</sup> Regressions (6) and (7) show that the inverted-U relationship survives the use of instruments for income. In these regressions with limited controls, the coefficients on the linear and quadratic origin income terms remain highly significant, and their respective signs continue to be consistent with an inverted-U shape. Our calculations yield that the 'emigration-maximizing' level of income according to this model and specification is around \$6,000, which implies that roughly 40% of countries in our sample would fall on the downward-sloping portion of the hump.

These initial results suggest the existence of a non-linear causal effect of income on emigration when controlling for destination income, population and distance alone. However, as more controls are added, the relationship loses statistical significance and eventually turns into a (statistically insignificant) largely upward sloping relationship.

<sup>&</sup>lt;sup>17</sup> We first ran a 2SLS regression including all controls in our data set. We then ran the same regression dropping the controls that were insignificant in the first and not included in our baseline model,

Table 1 column (7). This latter regression is reported in Table 1 column (10) and is the focus of our analysis (when all controls are included, 'Failed Govt., Coups (Origin)' is significant, but when limited controls are used, it becomes insignificant).

As we observe, the inclusion of three variables capture exogenous characteristics of the country-pair that account for the loss of significance of the effect of origin income. These variables are common border, common language, and former colonial relationship. The significance of these variables suggests that the apparent non-linear effect of income on migration is capturing the fact that middle income countries are also 'near' developed countries. These countries will therefore tend to have high emigration rates-because of their proximity to developed countries—and simultaneously be relatively developed, at least in comparison to other developing countries. In contrast, we also observe that the coefficients on destination income and destination income squared retain their significance with the inclusion of these additional controls in the fully-specified regression (10). This suggests that for most countries rising income leads to higher immigration levels. This latter result is consistent with the notion that as a country's income rises, it becomes more attractive for migrants from other countries. And while the coefficients of destination income in (10) appear to suggest that there exists an inverted U-shaped relationship between destination income and migration, all of the countries in our sample fall below the income threshold on the upward-sloping part of the curve in the full specification. This pattern suggests that while there may be a decrease in the returns of destination income to immigration, the relationship is positive for all countries.

Other factors associated with higher levels of development are also positively correlated with higher immigration levels and can be seen as acting as 'pull factors'. For example, the negative coefficient on the share of population above 65 suggests that a lower share of working-age population tends to be associated with higher immigration levels. Our results also show that immigration policies matter: places where policy makers are more receptive to higher levels of immigration tend to be more popular destinations. Looking at the standardized coefficients for our preferred 2SLS estimation in Table 3, we note that the ten variables that have the largest impact on the size of bilateral stocks pertain to the country of destination.

In general, we find that destination country variables matter a great deal, but we caution that the interpretation of many of these additional variables can vary with adjustments in the specification and the selection of instruments. In order to show that this is the case, the third panel of Table 1 reruns the same regressions with an additional set of variables in the instrument set, namely a set of colony dummies. The argument in favor of these instruments is that former colonial status is likely to be related to development today and is clearly exogenous. Excludability is more questionable, although it is to a great extent attenuated by the fact that colony-colonizer dummies are included in the set of explanatory variables. What is important to note is that in this alternative specification the signs of several of the explanatory variables change—such as life expectancy, education, and democracy. However, what is important for our key coefficients of interest is that the inverted-U relationship between income and migration disappears once controls are introduced (in fact, the specification with the most controls actually delivers a negative linear relationship).

## 4.2. The effect of migration on development

Now that we have explored how income levels have both push and pull effects, we examine the effect migration itself has on income. As Rodrik, et al (2004) show, when comparing three widely analyzed channels of growth—trade, institutions, and geography—while using instruments to control for endogeneity, the positive effect of institutions appears to dominate all other channels. Using a similar approach, we add migration into the equation—using geographical instruments—and observe whether migration plays a similar role in development. Felbermayr, Hiller, and Sala (2008) have performed a similar analysis finding a positive significant effect of migration on income, but include additional variables for market size and financial openness, whose exogeneity is questionable and fail to instrument for institutions.

In contrast to Felbermayr, et al (2008), we closely follow the approach pioneered by Frankel and Romer (1999) and refined by Rodrik, et al (2004) in which income is regressed

only on exogenous variables. This is achieved either by using variables which are naturally exogenous (e.g., geography) or by instrumenting for endogenous variables. In such a strategy, omitted variable bias is not a concern: as we are certain of the exogeneity of the explanatory variables, any correlation between them and an omitted variable would reflect causality from the former to the latter. However, introducing endogenous controls in the regressions, as Felbermayr, et al (2008) do, can seriously bias the coefficient estimates on the exogenous variables.<sup>18</sup>

In Table 4, we show the results of our estimations of GDP per capita on immigration, geography, institutions, and trade openness.<sup>19</sup> Immigration is instrumented by the predicted values of a regression of the log of immigrant stocks on several exogenous geographical variables: log of bilateral distance, a dummy variable for a shared border, absolute latitude of the origin country, and dummy variables for origin and destination landlocked countries. Rule of law measures institutional quality and is instrumented by ESM rates. Trade openness is instrumented by the geography-based Frankel-Romer openness index. In the restricted sample, we use only countries with available ESM rates data.<sup>20</sup> In the full sample, we set ESM rates equal to zero for former colonial powers and ex-Soviet countries. We focus our analysis on the full 2SLS estimations, (13)-(14).<sup>21</sup>

While immigration enters positive and highly significant in the simpler specifications—as found by Felbermayr, et al (2008)—it becomes insignificant whenever institutions are accounted for. In the full 2SLS estimation, the effect of immigration vanishes with the addition of rule of law. Moreover, in the fully specified model, (16), its magnitude is

<sup>&</sup>lt;sup>18</sup> See Wooldridge (2002).

<sup>&</sup>lt;sup>19</sup> In order to compare the magnitude of each variable's effect, we standardize all regressors by subtracting the mean and dividing by the standard deviation.

<sup>&</sup>lt;sup>20</sup> This sample is the same as that used by Rodrik, et al (2004).

<sup>&</sup>lt;sup>21</sup> In addition to increasing the total number of observations, the power of the instruments appears stronger in the full sample. First-stage regressions, reported in Table 6, show our instruments to be valid for all endogenous variables except immigration in the restricted sample, (4). We also observe that not all of the endogenous variables have the expected relationship to the corresponding instruments, particularly emigration, (3) and (7). It may be the case that the Frankel-Romer index, which is also based on geographic variables, is capturing part of the effect the other instruments.

very small, indicating that even if there is an effect, its impact is negligible compared to the other channels. In

Table 5 we find that emigration has an even less pronounced effect, while institutions continue to dominate, further confirming the results of Rodrik, et al (2004). In Table 5 we also include both immigration and emigration and find that their effect remains small and insignificant while institutions persist. Our results confirm the finding of Rodrik, et al (2004) that institutions trump other sources of growth after endogeneity is properly accounted for.

In sum, we fail to find any conclusive evidence suggesting that migration has a substantial effect on development, much less one that rivals that of institutions.

## **5.** Conclusions

Studies of the relationship between migration and development have traditionally had to deal with a set of distinct issues that hampered meaningful analysis. One is the need to properly account for the existence of a non-linear relationship between income and development, as suggested by many theoretical specifications. The second one is the need to take into account that migration decisions are affected by variables in both destination and origin countries and thus cannot be studied by the use of country aggregates that do not distinguish by both source and destination of migrants. A third issue is the need to study a process in which reasonable hypotheses about both directions of causation between development and income have been postulated.

This paper has addressed these three issues by presenting empirical estimates of a gravity model of bilateral migration that properly accounts for non-linearities and tackles causality issues through an instrumental variables approach. In contrast to previous contributions in this literature, which were limited to migration to OECD countries, we have estimated the model using a matrix of bilateral migration stocks for 127 countries (Migration DRC, 2007), allowing us to properly take into account migration to developing countries which accounts for 49 percent of international migration.

This exercise delivers a set of interesting results. First, the inverted-U relationship (or 'migration hump') between income at origin and income survives the more demanding bilateral stocks specification even after adding a large set of controls capturing conditions at home and destination as well as pair-specific variables. Second, although the relationship also survives in the basic IV estimation, it loses significance as we add more variables in the IV specification. Our two alternative IV specifications (which differ in the instrument list) deliver a similar message: the inverted-U does not survive the addition of controls for variables that capture the characteristics of destination economies or exogenous characteristics of the destination-origin pair.

It is best to illustrate these results with an example. Consider the cases of Morocco, Turkey and Mexico, all three countries with moderate levels of development. The high emigration rates of these three countries (9.0, 4.2 and 8.1 percent, respectively, as opposed to a world average of 3 percent) appear to confirm the migration 'hump' hypothesis whereby emigration rates increase and then decline with development. But these three countries also border highly developed regions (the United States and Europe). Furthermore, their high levels of development (in comparison to other countries in their respective regions) may arguably be caused by their proximity to developed countries. We can only distinguish between the effect of proximity and the effect of development if we can convincingly have exogenous sources of variation in development, which we obtain through our IV techniques. Once we do that, we find that proximity sweeps out the effect of development in both of our specifications. So while it is true that countries with middle levels of development have greater emigration rates, this appears to be caused by their proximity (geographical and cultural) to developed countries rather than the development process.

This paper has also studied the effect that migration has on both destination and origin country income by using geographically-induced differences across nations in immigration and emigration as sources of exogenous variations. In the spirit of Rodrik et al. (2004), we run a 'horse-race' between migration, institutions, trade and geography. While migration appears to have a significant effect in some of the simple specifications, its effect is

trumped by that of institutions, so that our preferred specification fails to find any significant effect of either immigration or emigration on income. This evidence supports the hypothesis that migration is best viewed from the standpoint of the scope that it offers to enhance *individual* opportunities rather than through its effect on aggregate economic performance.

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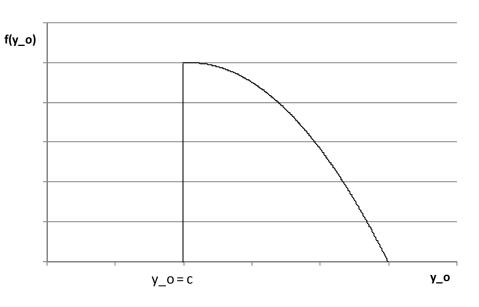
## **Data sources**

We merge data on the estimated stocks of international migrants with information on geographic, cultural, institutional, socio-economic, and demographic factors from origin and destination countries. Our source for data on international migrants is the University of Sussex/World Bank Global Migrant Origin Database. This consists of a 226x226 bilateral matrix of origin and destination stocks derived from the 2000 round of national population censuses. We use the fourth version, which combines place of birth and citizenship reporting mechanisms to create the first single, complete matrix of worldwide international migrant stocks. The fourth version is also the most up-to-date version available at this time (March 2007). Given the extensive list of independent variables required by our model (see below), we restrict the matrix size to 127x127 due to the limited data available for many of the smaller countries (i.e., New Caledonia and San Marino) as well as for irregular or politically complex countries (i.e., Afghanistan, North Korea, and West Bank and Gaza). Overall, this reduces the estimated sample size from 175 million to 166 million worldwide migrants (or 95% of the estimated total).

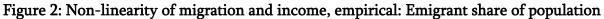
Data sources and summary statistics for all of the regressors used in the empirical model are documented in Appendices 1 and 2. Geographic and cultural information, including weighted distance between countries, common language, colonial ties, shared border, absolute latitude and whether a country is landlocked, comes from the *Centre d'études prospectives et d'informations internationales* (CEPII). Institutional measures come from a variety of sources: governance is based on the Rule of Law Index from Governance Matters VII, political instability is based on the Political Instability Task Force, democracy is based on the Polity IV Index and immigration policy is based on the World Population Policies from the United Nations. Regarding socio-economic data, GDP per capita and exports and services as a percentage of GDP are derived from the World Development Indicators; life expectancy and the education index are derived from the United Nations; and economic freedom is based on the Heritage Foundation's Index of Economic Freedom.

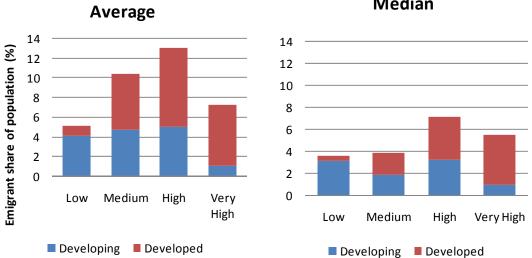
Finally, all demographic variables, including total population, the age structure of the population, population density and urban population, are derived from the United Nations Population Division. All independent variables correspond to the average value for the 1991-2000 period.

# **Tables and Figures**

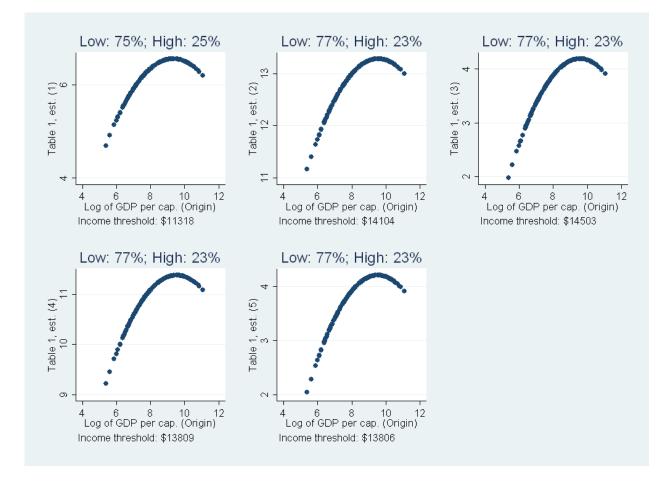


## Figure 1: Non-linearity of migration and income, theoretical





Median



## Figure 3: Emigration predicted by origin income, OLS

Ln(Stocks)			OLS					2SLS				2SLS with	former colon	y dummies	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)
Log of Distance	-1.316	-1.266	-1.016	-1.064	-1.053	-1.349	-1.3	-0.962	-0.893	-0.54	-1.31	-1.193	-0.905	-0.909	-1.318
	[0.0215]***	[0.0226]***	[0.0222]***	[0.0169]***	[0.0232]***	[0.0240]***	[0.0246]***	[0.0253]***	[0.0662]***	[0.194]***	[0.0230]***	[0.0260]***	[0.0262]***	[0.0632]***	[0.0443]***
Log of population (Origin)	0.589	0.648	0.627	0.596	0.63	0.61	0.666	0.633	0.618	0.587	0.586	0.657	0.638	0.634	0.742
o 11 ( o ,	[0.0115]***	[0.0109]***	[0.0103]***	[0.00627]***	[0.0100]***	[0.0128]***	[0.0120]***	[0.0113]***	[0.0316]***	[0.0404]***	[0.0121]***	[0.0128]***	[0.0126]***	[0.0318]***	[0.0349]***
Log of population (Dest.)		0.799	0.777	0.662	0.899		0.799	0.762	0.596	0.328		0.759	0.739	0.578	0.794
		[0.0108]***		[0.00619]***			[0.0118]***	[0.0115]***	[0.0443]***	[0.268]		[0.0135]***	[0.0131]***	[0.0333]***	[0.0330]***
Log of GDP per cap. (Origin)	2.22	2.317	2.39	1.778	2.393	7.044	6.696	1.086	-2.166	-0.88	1.293	1.042	0.658	0.439	-5.523
	[0.188]***	[0.181]***	[0.172]***	[0.119]***	[0.165]***	[1.053]***	[1.040]***	[0.979]	[2.628]	[3.188]	[0.736]*	[0.830]	[0.829]	[1.882]	[2.559]**
Log of GDP per cap. Sq. (Origin)	-0.119	-0.121	-0.125	-0.096	-0.125	-0.409	-0.384	-0.041	0.158	0.077	-0.0621	-0.0413	-0.0144	0.00153	0.218
tog of opt per capitod; (origin)	[0.0112]***			[0.00703]***		[0.0632]***	[0.0624]***	[0.0587]	[0.158]	[0.192]	[0.0439]	[0.0495]	[0.0494]	[0.112]	[0.154]
Log of GDP per cap. (Dest.)	[0.0112]	-5.579	-5.538	-4.507	-5.839	[0.0032]	-4.955	-10.12	-39.06	49.07	[0.0435]	-16.41	-16.99	-45.87	-37.87
Log of Obr per cap. (Dest.)		[0.169]***			[0.206]***		[0.885]***	[0.856]***	[7.338]***	[20.40]**		[0.635]***	[0.616]***	[4.214]***	[2.732]***
Log of GDP per cap. Sq. (Dest.)		0.362	0.359	0.294	0.357		0.32	0.636	1.773	-2.182		1.012	1.051	2.161	2.172
Log of GDP per cap. 3q. (Dest.)			• <b>[0.00969]</b> ***					[0.0514]***		[ <b>0.964</b> ]**		[0.0380]***			[0.163]***
Life Fundationary (Deat)		[0.0101]***	[0.00969]	0.0212	0.0272		[0.0532]***	[0.0514]	0.546	-0.503		[0.0380]	[0.0368]	0.591	0.178
Life Expectancy (Dest.)						ļ									
					[0.00368]***				[0.0988]***	[0.205]**					[0.0150]***
Population under 14 (%) (Dest.)				3.204	10.08				-19.35	39				-17.58	-6.702
				[0.212]***	[0.566]***				[4.505]***	[11.78]***				[4.442]***	[2.073]***
Urban Population (%) (Dest.)				0.823	0.631				22.01	-14.79				22.52	2.287
				[0.0768]***					[3.881]***	[6.123]**				[3.443]***	[0.387]***
Population over 65 (%) (Dest.)					17.01					51.58					-37.22
					[1.140]***					[26.29]**					[5.975]***
Services/GDP (Dest.)					2.205					-16.03					6.374
					[0.202]***					[7.837]**					[0.525]***
Exports/GDP (Dest.)					0.982					-7.082					1.736
					[0.0852]***					[3.605]**					[0.250]***
Polity IV (Dest.)					-0.0413					-0.3					0.0242
					[0.00409]***					[0.106]***					[0.00856]***
Education Index (Dest.)					1.539					-8.066					3.694
					[0.149]***					[3.719]**					[0.365]***
Log of population density (Dest.)					-0.215					1.25					-0.585
					[0.0135]***					[0.570]**					[0.0392]***
Immigration Policy (Dest.)					0.0435					1.28					0.073
<b>o</b> ,, , ,					[0.0390]					[0.530]**					[0.0734]
Common Border			1.99	1.601	1.858			2.198	1.83	3.295			2.411	1.848	1.324
			[0.126]***	[0.104]***	[0.123]***			[0.140]***	[0.304]***	[0.640]***			[0.152]***	[0.317]***	[0.196]***
Former Colony of Dest/Origin			1.971	2.164	1.913			1.766	1.826	1.88			1.597	1.83	1.971
, , , , , , , , , , , , , , , , , , , ,			[0.144]***	[0.114]***	[0.141]***			[0.153]***	[0.322]***	[0.354]***			[0.165]***	[0.331]***	[0.189]***
Common Language			1.276	1.105	1.238			1.263	2.101	0.724			1.234	2.068	1.387
				[0.0354]***				[0.0538]***		[0.327]**				[0.221]***	[0.0825]***
Failed Govt., Coups (Dest.)			[0:0010]	[0:000.1]	0.0979			[0:0000]	[0.215]	0.334			[0:00000]	[0.222]	-0.53
ranca Gove, coups (Dest.)	1				[ <b>0.0410</b> ]**					[ <b>0.176]</b> *					[0.0791]***
Failed Gover Coups (Origin)	1				0.0746					0.214					0.111
Failed Govt., Coups (Origin)	1														
Constant	2.804	2 214	0.00075	0.010	[0.0381]*	22.41	17.04	22.05	150	[0.144]	0.120	F1 00	F2 22	172.0	[0.0913]
Constant	-3.804	2.214	-0.00975	-0.918	-7.281	-23.41	-17.84	23.05	158	-220.7	-0.129	51.09	52.33	172.9	158.6
	[0.826]***	[1.092]**	[1.039]	[0.716]	[1.284]***	[4.309]***	[5.716]***	[5.399]***	[33.44]***	[81.45]***	[3.042]	[4.389]***	[4.339]***	[19.76]***	[16.23]***
<u>al</u>	00.757			00.075		00 FFF									
Observations	20,757	14,713	14,713	28,070	14,713	20,757	14,713	14,713	14,713	14,713	20,757	14,713	14,713	14,713	14,713
R-squared	0.231	0.498	0.546	0.532	0.578	0.206	0.476	0.518			0.23	0.365	0.39		

## Table 1: Bilateral migration equation estimates, OLS and 2SLS

Dependent variable is the natural log of bilateral stocks between countries. In columns (6)-(10), the 2SLS regressions use European Settler Mortality Rates and their squares for origin and destination countries as instruments. We first ran a 2SLS regression including all controls in our data set. We then ran the same regression dropping the controls that were insignificant in the first and not included in our baseline model, column (7). This latter regression is reported in column (10) and is the focus of our analysis (when all controls are included, 'Failed Govt., Coups (Origin) is significant, but when limited controls are used, it becomes insignificant). We then perform these same regressions using dummies for origin and destination for former colonies in columns (11)-(15). First-stage regressions reported in Table 2. Robust standard errors in brackets. Significance indicated by \*\* p<0.01, \*\* p<0.05, \* p<0.1.

## Table 2: Bilateral migration equation estimates, 2SLS first stage regressions

	Origin	Origin		Dest.	Origin	Origin		Dest.
	income	income sq.	Dest. income	income sq.	income	income sq.	Dest. income	income sq.
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Origin Log of ESM (Origin)	0.0147	-0.0911	0.00287	0.0289	-0.0674	-1.562	-0.00267	-0.0914
	[0.00714]**	[0.113]	[0.00451]	[0.0747]	[0.0108]***	[0.181]***	[0.00474]	[0.0812]
Origin Log of ESM Sq. (Origin)	-0.0315	-0.439	-0.000491	-0.00549	-0.0427	-0.639	0.000232	0.00997
	[0.00120]***	[0.0188]***	[0.000730]	[0.0121]	[0.00184]***	[0.0303]***	[0.000770]	[0.0133]
Destination Log of ESM (Dest.)	-0.00715	-0.108	0.0152	0.446	-0.0268	-0.459	0.0131	0.401
	[0.00922]	[0.148]	[0.00655]**	[0.104]***	[0.0127]**	[0.215]**	[0.00668]*	[0.108]***
Destination Log of ESM Sq. (Dest.)	0.000517	0.00787	-0.0051	-0.121	0.00221	0.0383	-0.00407	-0.0984
о т <i>с</i> ,	[0.00142]	[0.0229]	[0.00102]***	[0.0161]***	[0.00197]	[0.0335]	[0.000988]***	
Origin Former Colony (Origin)								
Destination Former Colony (Dest.)								
Rule of Law (Origin)	0.759	13.59	-0.00132	-0.0194				
	[0.00474]***	[0.0787]***	[0.00387]	[0.0644]				
Rule of Law (Dest.)	-0.000402	-0.00355	0.277	5.956				
	[0.00983]	[0.158]	[0.00555]***	[0.0948]***				
.og of Distance	0.0226	0.314	-0.0177	-0.221	0.0564	0.919	-0.00731	0.00329
	[0.00707]***	[ <b>0.115]</b> ***	[ <b>0.00471</b> ]***	[0.0792]***	[0.00976]***	[ <b>0.167</b> ]***	[0.00496]	[0.0864]
Common Border	0.0645	0.971	-0.0551	-0.735	-0.0931	-1.852	-0.0815	-1.304
common border	[ <b>0.0364</b> ]*	[ <b>0.578</b> ]*	[ <b>0.0207]</b> ***	[ <b>0.343]</b> **	[ <b>0.0480</b> ]*	[0.806]**	[ <b>0.0219</b> ]***	[ <b>0.376</b> ]***
Common Languago	0.0627	0.979	0.0333	0.506	0.309	5.387	0.0777	1.466
Common Language	[0.0169]***	[ <b>0.275</b> ]***	[ <b>0.0110</b> ]***	[ <b>0.181]</b> ***	[ <b>0.0215</b> ]***	[0.363]***	[ <b>0.0116</b> ]***	[0.198]***
formar Calany of Dact/Origin		0.751	-0.0333			2.9	-0.0337	-0.647
ormer Colony of Dest/Origin	0.0446			-0.64	0.165	2.9 [0.807]***		
inited Courts (Oninin)	[0.0338]	[0.546]	[0.0210]	[0.356]*	[0.0467]***		[0.0225]	[0.393]*
ailed Govt., Coups (Origin)	-0.236	-3.199	-0.00247	-0.0327	-0.923	-15.51	0.000554	0.0243
	[0.0128]***	[0.202]***	[0.00744]	[0.124]	[0.0146]***	[0.238]***	[0.00695]	[0.120]
og of population (Origin)	-0.000825	0.0941	0.000716	0.0106	-0.0201	-0.251	0.000897	0.0143
	[0.00376]	[0.0628]	[0.00196]	[0.0327]	[0.00469]***	[0.0806]***	[0.00208]	[0.0358]
ailed Govt., Coups (Dest.)	-0.00285	-0.0451	0.0461	1.06	-0.00721	-0.124	-0.0153	-0.262
	[0.0129]	[0.209]	[0.00783]***	[0.127]***	[0.0177]	[0.302]	[0.00761]**	[0.123]**
Polity IV (Dest.)	-0.000323	-0.00449	0.0122	0.172	-0.00129	-0.0218	0.0204	0.348
	[0.00117]	[0.0189]	[0.000829]***	[0.0133]***	[0.00157]	[0.0268]	[0.000801]***	
ife Expectancy (Dest.)	8.06E-05	0.00162	0.0265	0.354	0.000313	0.00589	0.0342	0.522
	[0.00116]	[0.0187]	[0.000995]***	[0.0160]***	[0.00156]	[0.0265]	[0.00109]***	[0.0184]**
ducation Index (Dest.)	-0.0364	-0.539	0.691	11.26	-0.0956	-1.603	0.44	5.846
	[0.0478]	[0.769]	[0.0334]***	[0.535]***	[0.0646]	[1.097]	[0.0358]***	[0.595]***
Exports/GDP (Dest.)	-0.00629	-0.105	0.665	11.4	-0.0197	-0.343	0.821	14.76
	[0.0247]	[0.398]	[0.0156]***	[0.250]***	[0.0333]	[0.566]	[0.0160]***	[0.273]***
ervices/GDP (Dest.)	-0.0227	-0.341	1.114	16.65	-0.112	-1.931	1.728	29.87
	[0.0663]	[1.067]	[0.0508]***	[0.855]***	[0.0870]	[1.479]	[0.0537]***	[0.926]***
og of population (Dest.)	-0.00167	-0.0253	0.0635	1.131	-0.00323	-0.0531	0.0693	1.255
	[0.00369]	[0.0595]	[0.00247]***	[0.0423]***	[0.00512]	[0.0870]	[0.00261]***	[0.0463]**
Population under 14 (%) (Dest.)	-0.0401	-0.742	0.494	20.54	-0.202	-3.648	0.592	22.65
	[0.187]	[3.015]	[0.110]***	[1.750]***	[0.260]	[4.428]	[0.114]***	[1.903]***
Population over 65 (%) (Dest.)	0.0316	0.0272	3.863	97.8	-0.125	-2.739	6.705	159
	[0.376]	[6.057]	[0.225]***	[3.746]***	[0.505]	[8.581]	[0.221]***	[3.811]***
og of population density (Dest.)	0.000333	0.00372	-0.0764	-1.113	-0.00251	-0.0474	-0.0844	-1.285
	[0.00409]	[0.0659]	[0.00246]***	[0.0398]***	[0.00566]	[0.0961]	[0.00265]***	[0.0454]***
Jrban Population (%) (Dest.)	-0.0124	-0.21	1.229	20.52	-0.0232	-0.405	1.172	19.27
	[0.0369]	[0.594]	[0.0329]***	[0.558]***	[0.0512]	[0.870]	[0.0343]***	[0.592]***
mmigration Policy (Dest.)	-0.00262	-0.0337	-0.132	-2.292	-0.0122	-0.205	-0.117	-1.973
	[0.0116]	[0.188]	[0.00784]***	[0.130]***	[0.0161]	[0.274]	[0.00877]***	[0.152]***
Constant	8.856	78.07	3.615	-9.509	9.629	[0.274] 91.9	2.513	-33.23
onstant	0.000 [0.171]***	78.07 [ <b>2.765]</b> ***	5.015 [ <b>0.104]</b> ***	-9.509 [ <b>1.720]</b> ***	9.629 [ <b>0.228]</b> ***	[ <b>3.888</b> ]***	2.515 [0.105]***	-33.23 [ <b>1.827</b> ]***
Observations	16002	16002	16002	16002	16002	16002	16002	16002
	0.774	0.793						
R-squared			0.911	0.913	0.562	0.552	0.9	0.895
test: ESM variables=0	1494	1397	12.25	20.14	2934	2862	8.136	14.71
Prob>F	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00

Table 3: Gravity model stand	dardized coefficients
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	OLS			2SLS						
Туре	Variable	Effect on Emigration Rate	Туре	Variable	Effect on Emigration Rate					
Destination	GDP per capita, above	0.66	Destination	GDP per capita, below	16.08					
Destination	Total Population	0.47	Destination	Life Expectancy	-1.85					
Origin	Total Population	0.39	Destination	Population 0-14 (%)	1.46					
Destination	GDP per capita, below	-0.34	Destination	Urban Population (%)	-1.18					
Relationship	Weighted-distance	-0.29	Destination	Population 65+ (%)	0.84					
Origin	GDP per capita, below	0.27	Destination	Services (% of GDP)	-0.75					
Destination	Population 0-14 (%)	0.19	Destination	Exports (% of GDP)	-0.68					
Relationship	Common Language	0.14	Destination	Polity IV Index	-0.62					
Destination	Services (% of GDP)	0.12	Destination	Education Index	-0.62					
Destination	Polity IV Index	-0.11	Destination	Population density	0.62					
Destination	Education Index	0.09	Origin	GDP per capita, above	0.52					
Destination	Population 65+ (%)	0.08	Origin	Total Population	0.31					
Relationship	Colony	0.08	Destination	Immigration Policy	0.20					
Relationship	Border	0.08	Relationship	Border	0.18					
Destination	Life Expectancy	0.07	Destination	Total Population	0.18					
Origin	GDP per capita, above	-0.06	Relationship	Weighted-distance	-0.16					
Destination	Population density	-0.05	Relationship	Colony	0.08					
Destination	Urban Population (%)	0.04	Relationship	Common Language	0.08					
Destination	Immigration Policy	-0.03	Destination	War and Govt. Failure	0.05					
Destination	Exports (% of GDP)	0.02	Origin	War and Govt. Failure	0.03					
Origin	War and Govt. Failure	0.01	Origin	GDP per capita, below	0.00					
Destination	War and Govt. Failure	0.00	Destination	GDP per capita, above						

Reported are the standardized coefficients corresponding to regressions (5) and (10) in Table 1 and listed in descending order of absolute magnitude. Standardized coefficients for income are calculated separately for countries on the upward and downward sloping parts of the curve. In the 2SLS specification, all countries fall below the destination income threshold and are on the upward sloping part of the curve, thus there is no value for 'Destination GDP per capita, above'.

#### Table 4: Migration and income

Ln(GDP per capita)				Restricted s	ample				Full sample							
	OLS					2SLS				0	LS		2SLS			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
Immigration	0.727	0.634	0.275	0.217	0.64	0.953	-1.6	-0.0274	0.7	0.451	0.142	0.11	0.715	0.503	-0.859	-0.00682
	[0.138]***	[0.134]***	[0.142]*	[0.140]	[0.577]	[0.595]	[1.977]	[2.644]	[0.0969]***	[0.0921]***	[0.0807]*	[0.0790]	[0.187]***	[0.167]***	[0.828]	[0.791]
Distance from the Equator		0.523	0.227	0.337		0.478	-0.565	-0.667		0.477	0.216	0.214		0.612	-0.747	-0.867
		[0.115]***	[0.128]*	[0.132]**		[0.125]***	[0.479]	[0.463]		[0.0646]***	[0.0578]***	[0.0571]***		[0.0615]***	[0.624]	[0.641]
Rule of law			0.499	0.43			2.022	1.704			0.599	0.599			2.236	2.209
			[0.107]***	[0.109]***			[0.993]**	[0.957]*			[0.0534]***	[0.0532]***			[0.991]**	[0.968]**
Trade openness				0.163				-0.309				0.0568				-0.445
				[0.0873]*				[0.318]				[0.0545]				[0.386]
Constant	-0.155	0.128	0.00352	0.084	-0.178	0.186	-0.518	-0.35	0.111	0.0322	0.0123	0.016	0.157	0.175	-0.328	-0.314
	[0.1000]	[0.0841]	[0.0800]	[0.0826]	[0.169]	[0.147]	[0.475]	[0.521]	[0.0791]	[0.0697]	[0.0540]	[0.0533]	[0.0872]*	[0.0636]***	[0.270]	[0.260]
Observations	79	79	79	79	79	79	79	79	125	125	125	125	107	107	107	107
R-squared	0.229	0.375	0.504	0.53					0.213	0.461	0.691	0.694				

Dependent variable is the natural log of GDP per capita. Immigration is measured as the immigrant share of population. ESMR, the Frankel-Romer openness index, and the geographic component of immigration are used as instruments for rule of law, trade openness, and immigration, respectively. See Table 6 for first-stage regressions. Robust standard errors in brackets. Significance indicated by \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

### Table 5: Migration and income, additional specifications

Ln(GDP per capita)				Restricted	sample							Full sam	ple			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
Emigration	0.186	0.277	-0.329	0.0106	-0.544	1.152	0.386	-1.889	0.724	0.391	-0.725	0.00805	1.433	-0.0622	-0.572	0.0588
	[0.156]	[0.122]**	[0.410]	[0.530]	[2.045]	[13.81]	[8.292]	[44.54]	[0.489]	[0.217]*	[0.744]	[0.958]	[1.137]	[0.164]	[0.583]	[0.798]
Immigration					2.67	-2.965	-3.503	9.971					-0.812	0.574	-0.207	-0.0507
					[7.709]	[46.73]	[40.24]	[246.5]					[2.016]	[0.209]***	[0.569]	[0.488]
Distance from the Equator		0.619	-0.526	-0.667		1.056	-0.613	-0.473		0.698	-0.877	-0.865		0.6	-0.855	-0.859
		[0.142]***	[0.460]	[0.452]		[6.818]	[1.392]	[5.974]		[0.0670]***	[0.834]	[0.685]		[0.0610]***	[0.752]	[0.648]
Rule of law			1.631	1.694			2.491	-0.729			2.209	2.204			2.228	2.209
			[0.554]***	[0.511]***			[10.08]	[61.48]			[1.124]*	[0.939]**			[1.102]**	[0.959]**
Trade openness				-0.316				-0.442				-0.449				-0.457
				[0.306]				[4.735]				[0.506]				[0.481]
Constant	-0.317	0.062	-0.239	-0.346	0.269	-0.321	-0.851	1.352	0.157	0.163	-0.314	-0.312	0.138	0.175	-0.323	-0.314
	[0.0966]***	[0.101]	[0.206]	[0.234]	[1.705]	[5.982]	[6.981]	[41.84]	[0.145]	[0.0790]**	[0.275]	[0.248]	[0.293]	[0.0689]**	[0.275]	[0.258]
Observations	79	79	79	79	79	79	79	79	107	107	107	107	107	107	107	107

Two-stage least squares regressions. Dependent variable is the natural log of GDP per capita. Immigration and emigration are measured as the immigrant and emigrant share of population, respectively. ESMR, the Frankel-Romer openness index, and the geographic component of immigration and emigration are used as instruments for rule of law, trade openness, and emigration, respectively. See Table 6 for first-stage regressions. Robust standard errors in brackets. Significance indicated by \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

		Restricte	d sample			Full s	ample	
	Rule of Law	Trade openness	Emigration	Immigration	Rule of Law	Trade openness	Emigration	Immigration
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
ESMR	-0.327	-0.261	-0.113	-0.105	-0.3	-0.119	-0.168	-0.132
	[0.113]***	[0.0993]**	[0.0903]	[0.102]	[0.157]*	[0.148]	[0.149]	[0.118]
Frankel-Romer	0.129	0.8	0.178	0.0753	0.116	0.707	0.169	0.0767
	[0.125]	[0.117]***	[0.0856]**	[0.143]	[0.125]	[0.106]***	[0.0774]**	[0.0973]
Immigration, geo	2.333	2.103	10.93	2.477	0.323	-0.132	-0.52	1.236
	[3.364]	[1.649]	[4.806]**	[3.716]	[0.453]	[0.313]	[0.887]	[0.326]***
Emigration, geo	-2.12	-2.026	-9.929	-2.24	-0.242	0.15	0.836	-0.945
	[3.228]	[1.570]	[4.537]**	[3.570]	[0.502]	[0.289]	[0.769]	[0.364]**
Distance from Equator	0.548	-0.143	0.173	0.142	0.455	-0.134	-0.177	0.0649
	[0.175]***	[0.148]	[0.0891]*	[0.108]	[0.147]***	[0.139]	[0.154]	[0.0981]
Constant	0.122	-0.13	0.21	-0.105	0.195	-0.122	-0.0349	-0.0488
	[0.150]	[0.108]	[0.127]	[0.121]	[0.0848]**	[0.0742]	[0.106]	[0.0629]
Observations	79	79	79	79	108	107	108	108
R-squared	0.413	0.597	0.49	0.116	0.503	0.525	0.227	0.249
F-test	10.16	25	5.946	1.464	30.73	29.06	5.712	24.3
Prob>F	0.00	0.00	0.00	0.21	0.00	0.00	0.00	0.00

## Table 6: Migration and income, 2SLS first stage regressions

Robust standard errors in brackets. Significance indicated by \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

## Table A1: Sources

Dependent Variables	Source	Units and coding						
Log of migrant stocks in world	University of Sussex/World Bank Global Migrant Origin Database, 2000	The reference period is the 2000 round of population censuses						
Log of inflow of foreign population in OECD country	OECD International Migration Database, 1996- 2005 average	Foreign population is defined by nationality						
Log of stock of foreign population in	OECD International Migration Database, 1996-	Foreign population is defined by nationality						
OECD country	2005 average							
Independent Variables	Source	Notes						
Country Pair								
Log of weighted distance between countries	Centre D'Etudes Prospectives et D'Informations Internationales (CEPII)	Bilateral distances between biggest cities weighted by population						
Common language	CEPII	1 if language is spoken by at least 9% of the population in both countries						
Colony	CEPII	1 if long period of colonial relationship with substantial participation in the governance of the colonized country						
Shared border	CEPII	1 if shared border						
Geographic								
Absolute latitude	CEPII							
Landlocked	CEPII	1 if country is landlocked						
Institutional								
Governance: Rule of law index	Governance Matters VII: Aggregate and Individual Governance Indicators, 1996-07	-2.5 to 2.5 where 2.5 = best governance						
Political instability	Political Instability Task Force, Historical State Armed Conflicts and Regime Crises, 1955-2007	1 if country experienced war, genocide, politicide and/or adverse regime change duri time period						
Democracy: Polity IV index	Polity IV Project: Political Regime Characteristics and Transitions, 1800-2007	-10 (hereditary monarchy) to +10 (consolidated democracy)						
Policy on immigration level	World Population Policies 2007, United Nations Population Division	-1 (lower), 0 (no intervention or maintain), +1 (raise)						
Socio-economic								
Log of GDP per capita	World Development Indicators (WDI)	Constant PPP-adjusted 2005 international \$						
Life expectancy at birth	World Population Prospects: 2006 Revision, United Nations Population Division	Total years						
Education index	United Nations Development Programme, Human Development Report Office	Combination of literacy and gross enrolment ratio. 0-1 where 1 is 100% literacy and enrolment						
Exports of goods and services	WDI	% of GDP						
Services, etc., value added	WDI	% of GDP						
Economic freedom	Heritage Foundation's Index of Economic Freedom	0-100 where 100 represents maximum freedom						
Demographic								
Log of total population	World Population Prospects: 2006 Revision, United Nations Population Division							
Population under 15	Ibid	% of total						
Population over 64	Ibid	% of total						
Population density	Ibid for total population; CEPII for country area	Sq. km.						
Urban population	World Urbanization Prospects: 2007 Revis-ion, United Nations Population Division	% of total						
Additional Information								
Countries not included in the sample	Andorra, Afghanistan, Netherlands Antilles, Amer	rican Samoa, The Bahamas, Bermuda, Channel						
due to lack of data	Islands, Cayman Islands, Faeroe Islands, Microne Monaco, Marshall Islands, Northern Mariana Isla Democratic Republic of North Korea, French Poly Islands, and West Bank and Gaza	nds, Mayotte, New Caledonia, Palau, Puerto Rico,						
Methodology to interpolate missing data for countries in the sample	nodology to interpolate missing Missing values were calculated by taking the average value of the variable for countries with							

## Table A2: Summary statistics

Variable	Obs	Mean	Std. Dev.	Min	Max
Bilateral stocks	32,580	5,103	78,933	0	9,336,719
Absolute latitude	181	25.90	16.73	0.23	64.15
Landlocked	181	0.19	0.40	0	1
Rule of Law	181	-0.08	0.97	-2.20	2.02
Polity IV Index	181	2.57	6.20	-10	10
War and Govt. Failure	181	0.28	0.45	0	1
Immigration Policy	181	-0.29	0.45	-1	1
Ln GDP per capita	181	8.38	1.29	5.38	11.08
Life Expectancy	181	65.4	10.3	29.8	80.0
Education Index	181	0.74	0.20	0.12	0.99
Exports (% of GDP)	181	38.2%	25.8%	1.1%	212.3%
Services (% of GDP)	181	51.7%	13.5%	14.8%	83.9%
Total Population	181	15.54	1.95	10.63	20.91
Population 0-14 (%)	181	32.8%	10.5%	2.8%	49.7%
Population 65+ (%)	181	6.5%	4.3%	1.0%	17.4%
Urban population	181	51.3%	23.3%	7.3%	100.0%
Population density	181	4.02	1.51	0.38	8.65
ESMR	127	2.90	2.47	0.00	7.99